

9/12/16

Discrete Random Variables:

Ω : finite or countably infinite

$$X: \Omega \rightarrow \mathbb{R}$$

$$\Omega = \{\omega_1, \omega_2, \omega_3, \dots\}$$

Continuous Random Variables:

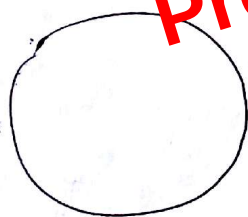
An r.v. X defined on an uncountably infinite sample space is called continuous if

$$P(X=x) = 0 \quad \forall x \in \mathbb{R}$$

$\Omega = [0, 1]$, cannot enumerate in a list
↖ all infinitely long sequences of 0's & 1's

Ex: $X \sim \text{uniform}(0, 1)$ if $\forall 0 \leq a < b \leq 1$
 $P(X \in [a, b]) = b - a$

Statistics:



population of individuals

N : size of population

N_D : # democrats

$N(1-p)$: # republicans

p : parameter

$$p = \frac{N_D}{N} \leftarrow \# \text{ democrats}$$

Statisticians provide an estimate of p based

on data $\rightarrow$ representative subset of the population
Data are random-generated through a random experiment.

Gathering data: SRSWR or SRSWOR

↑
with replacement

↑
without replacement

Y : response variable

$(x_1, \dots, x_m)$: covariates (the other variable)

Y : phenotype / response to certain kinds of treatment

Genotype: genetic information forms some of the covariates, environmental / cultural factors

Bivariate Density Function:

Let (X, Y) be a continuous random vector
i.e. $P(X=x, Y=y) = 0$

(X, Y) has a density function if there exists a non-negative function

$f: \mathbb{R}^2 \rightarrow [0, \infty)$ such that for any $A \subseteq \mathbb{R}^2$,
 $P((X, Y) \in A) = \int_A f(x, y) dx dy$

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$$= \int_{-\infty}^{\infty} \frac{1}{\pi} \mathbb{I} \left\{ y \in \left(-\sqrt{1-x^2}, \sqrt{1-x^2} \right) \right\} dy$$

$$= \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{1}{\pi} dy = \frac{y}{\pi} \Big|_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}}$$

(not uniform)

$$= \frac{2\sqrt{1-x^2}}{\pi}, \quad x \in [-1, 1]$$

$$f_Y(y) = \frac{2\sqrt{1-y^2}}{\pi} \mathbb{I}(y \in [-1, 1])$$

conditionals:

$$f_{Y|X=X}(y) = \frac{f(X,Y)}{f_X(x)} = \frac{\frac{1}{\pi} \mathbb{I}(x^2+y^2 \leq 1)}{\frac{2\sqrt{1-x^2}}{\pi}}$$

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$$= \frac{1}{2\sqrt{1-x^2}} \mathbb{I}(y \in (-\sqrt{1-x^2}, \sqrt{1-x^2}))$$

(uniform) $-1 \leq x \leq 1$

Given r.v. X

2 important functionals $f(x)$ satisfies are

$E(X)$ & $\text{var}(X)$

same as $\sum y_j P_{xy}(y_j)$

similar results if x continuous

x has pdf f_x

$$Y = g(x)$$

$$E(Y) = \int g(x) f_x(x) dx$$

Joint Distributions:

$(x_1, x_2, \dots, x_n)$ is a continuous random vector with pdf $f(x_1, x_2, \dots, x_n)$

$$P(x_1 \in (a_1, b_1), x_2 \in (a_2, b_2), \dots, x_n \in (a_n, b_n))$$

$$= \int_{a_1}^{b_1} \int_{a_2}^{b_2} \dots \int_{a_n}^{b_n} f(x_1, \dots, x_n) dx_n \dots dx_1$$

Think about $g(x_1, x_2, \dots, x_n)$

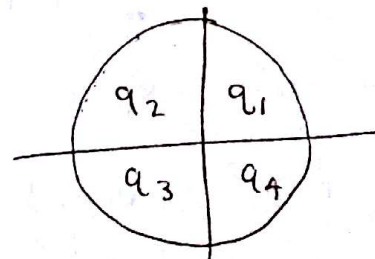
$$E(g(x_1, \dots, x_n)) = \int g(x_1, \dots, x_n) f(x_1, \dots, x_n) dx_1 \dots dx_n$$

(x, y)

$$f(x, y) = \frac{1}{\pi} \mathbb{I}\{x^2 + y^2 \leq 1\} \quad g(x, y) = xy$$

$$E(g(x, y)) = \int_{x^2 + y^2 \leq 1} xy \cdot \frac{1}{\pi} \mathbb{I}\{x^2 + y^2 \leq 1\} dx dy$$

$$= \int_{a_1} x y dx dy + \int_{a_2} x y dx dy + \int_{a_3} x y dx dy + \int_{a_4} x y dx dy = 0$$



We define $E[Y|X]$, the random variable as $\hat{Y}(x)$
 ↪ not the same as scalar $\hat{Y}$ from before

Properties of $E[Y|X]$:

$$(1) E(E(Y|X)) = E(Y)$$

$$(2) \text{var}(Y) = E(\text{var}(Y|X)) + \text{var}(E(Y|X))$$

$$\text{var}(Y|X=x) = \begin{cases} \sum (Y - E(Y|X=x))^2 P_{Y|X=x}(Y) & \text{discrete} \\ \int (Y - \hat{Y}(x))^2 f_{Y|X=x}(Y) dY & \text{continuous} \end{cases}$$

$$\hat{Y}(x) = E[Y|X=x]$$

$$\text{prediction: } E[(Y - \Psi(x))^2]$$

$$= E[(Y - E(Y|X) + E(Y|X) - \Psi(x))^2]$$

$$= E[(Y - E(Y|X))^2 + (E(Y|X) - \Psi(x))^2 +$$

$$2E[(Y - \hat{Y}(x))(\hat{Y}(x) - \Psi(x))]]$$

||
W

$$(X, W) \quad E(W) = E[E(W|X)]$$

$$E[W|X] = E[(Y - \hat{Y}(x))(\hat{Y}(x) - \Psi(x)) | X]$$

$$= (\hat{Y}(x) - \Psi(x)) E[(Y - E[Y|X]) | X]$$

$$= (\hat{Y}(x) - \Psi(x)) (E[Y|X] - E[Y|X]) = 0$$

$$= E[(Y - E[Y|X])^2] + E[E(Y|X) - \Psi(x)]^2$$

The best $\Psi(x)$ is $E[Y|X]$