

## Discrete Choice Model

- ▶ Utility of consumer  $i$  for product  $j$

$$u_{ij} = U(x_j, p_j, v_i)$$

- ▶  $x_j$  is a vector of product characteristics
- ▶  $p_j$  is the price of the product and
- ▶  $v_i$  is a vector of consumer characteristics

- ▶ Horizontal Model Example: Hotelling with quadratic costs

$$u_{ij} = \bar{u} - p_j - (x_j - v_i)^2$$

- ▶  $x_j$  is the location of the product along the line
- ▶  $v_i$  is the location of the consumer

## Random Utility Model (RUM)(McFadden)

- ▶  $U_{ij}^*$  usually specified as a sum of two parts

$$U_{ij}^*(x_j, p_j, p_z, y_i) = V_{ij}(x_j, p_j, p_z, y_i) + \varepsilon_{ij}$$

- ▶  $\varepsilon_{ij}$  i.i.d. across products and consumers; represents consumer tastes (observed by consumer but not by the researcher)
- ▶ What does it mean for tastes to be represented by product and consumer specific random terms?
  - ▶ product chosen is random from the researchers point of view
  - ▶ McFadden won the Nobel Prize for this in 2000
  - ▶ Assumptions about distribution of the  $\varepsilon_{ij}$ 's determines choice probabilities
- ▶ The probability that consumer  $i$  buys product  $j$  is

$$D_{ij}(p_1, \dots, p_j, p_z, y_i) = \Pr ob \{ \varepsilon_{i0}, \dots, \varepsilon_{ij} : U_{ij}^* > U_{ik}^*, \text{ for } j \neq k \}$$

## Independence of Irrelevant Alternatives (IIA)

- ▶ ratio of choice prob (odds ratio) does not depend on the number of alternatives available

$$\frac{s_{ij}}{s_{in'}} = \frac{\exp(V_{ij})}{\exp(V_{in'})}$$

- ▶ Red bus/blue bus problem: Walk or take red bus
  - ▶ If consumer walks half the time then  $s_{iW} = s_{iRB} = 0.5$
  - ▶ odds ratio walk/RB=1
- ▶ Introduce a red bus
  - ▶ odds ratio between walk/BB is 1
- ▶ But buses are perfect substitutes
  - ▶ new choice prob should be  $s_{iW} = 0.5$ ;  $s_{iRB} = s_{iBB} = 0.25$
  - ▶ new odds ratio should be walk/RB=2
- ▶ IIA is especially troubling if want to predict penetration of new products

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## Counter-intuitive substitution patterns:

- ▶ Not only from the distributional logit assumption
- ▶ Due to assumption that the only variance in consumer tastes comes through the i.i.d. product-specific terms  $\varepsilon_{ij}$
- ▶ Since i.i.d., there is no source of correlation in consumer tastes across similar products
- ▶ Changes to allow for more intuitive substitution patterns
  - ▶ Generalized EV models (GEV, Nested logit)
  - ▶ Mixtures of logits (K types of logit parameters)
  - ▶ Product differentiation model (Bresnahan, Stern, Trajtenberg 1997)
  - ▶ Random Coefficients Model of Demand (Berry, Levinsohn, and Pakes)

## Nested Logit Model

- ▶ Within-group correlation parameter is  $\sigma_g$
- ▶ Across nests, parameter  $\sigma$  (within  $(0,1)$ ) describes correlation between nests

$$u_{ij} = x_j\beta - \alpha p_j + \sigma_g v_{ig} + \varepsilon_{ij}$$

- ▶ Define the inclusive value of nest  $g$  as:

$$s_{ig} = \sum_{j \in g} \exp\left(\frac{u_{ij}}{1 - \sigma}\right)$$

- ▶ McFadden (1978) showed nested structure is consistent with RUM maximization iff the coefficients of the inclusive value lie within the unit interval
- ▶ More complicated forms of cross-product correlation in tastes do not lead to closed form expressions for shares (like Nested Logit does)

▶ need to compute a high dimensional integral and this is tough

▶ simulation methods help here

## Inversion Example: Berry Logit

- ▶ Simple example of the inversion step: MNL shares

$$\widehat{s}_{jt}(\delta) = \frac{\exp(\delta_{jt})}{1 + \sum_{k=1}^J \exp(\delta_{kt})}$$

$$\log \widehat{s}_{jt} = \delta_{jt} - \log \left( 1 + \sum_{k=1}^J \exp(\delta_{kt}) \right)$$

- ▶ but notice for outside good

$$\log \widehat{s}_{0t} = 0 - \log \left( 1 + \sum_{k=1}^J \exp(\delta_{kt}) \right)$$

- ▶ so  $\delta_{jt} = \log \widehat{s}_{jt} - \log \widehat{s}_{0t}$

- ▶ This implies

$$(\log S_{jt} - \log S_{0t}) = \delta_{jt} = X_{jt}\beta - \alpha p_{jt} + \xi_{jt}$$

- ▶ where  $S_{0t}$  is the share of the outside good
- ▶ Can be estimated by OLS
  - ▶ dependent variable  $\log S_{jt} - \log S_{0t}$
  - ▶ covariates  $X_{jt}, p_{jt}$ , and error term:  $\xi_{jt}$
- ▶ When there is not a closed form solution for the market share then solve for  $\xi$  structural error and construct moment condition
- ▶ Restrict the model predictions for product  $j$ 's market share to match the observed market shares

$$S_t^{obs} - s_t(\delta, \theta) = 0$$

- ▶ then solve for the demand side unobservable

$$\xi_{jt} = \delta_{jt}(S, \theta) - X_{jt}'\beta - \alpha p_{jt}$$



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## What are appropriate instruments?

- ▶ IV for  $j$  should be correlated with  $p_j$  but not with structural error  $\xi_j$
- ▶ Usual demand case: cost shifters
  - ▶ but we have cross-sectional (across products) data, so we require IV to vary across products within a market
- ▶ Example: cars, one natural cost shifter are wages in Michigan
- ▶ Here doesn't work because its the same across all products
  - ▶ if ran 2SLS with wages in Michigan as IV, first stage regression of price on wage would yield the same predicted price for all products

- ▶ One commonly used specification is the logit model with random (normal) coefficients

$$U_{ij} = X_j \beta_i - \alpha p_j + \xi_j + \varepsilon_{ij}$$

- ▶ The  $K$  random coefficients (one for each product characteristic) are

$$\beta_{ik} = \beta_k + \sigma_k v_{ik}$$

$$v_{ik} \sim N(0, 1), \text{ iid}$$

- ▶ it is useful to decompose utility into two parts

$$\mu_{ij} = \sum_k \sigma_k X_{jk} v_{ik}$$

$$\delta_j = X_j \beta_k - \alpha p_j + \xi_j$$

- ▶ So we can rewrite indirect utility as

$$u_{ij} = \delta_j + \mu_{ij} + \varepsilon_{ij}$$

## Intuition of the Estimation Algorithm

- ▶ The model is one of individual behavior, yet only aggregate data is observed.
- ▶ We can still estimate the parameters that govern the distribution of individuals
  - ▶ compute predicted individual behavior and aggregate over individuals, for a given value of the parameters,
  - ▶ obtain predicted market shares
- ▶ We then choose the values of the parameters that minimize the distance between these predicted shares and the actual observed shares
- ▶ The metric under which this distance is minimized is not the straightforward sum of least squares
- ▶ rather it is the metric defined by the instrumental variables and the GMM objective function
- ▶ It is this last step that somewhat complicates the estimation procedure

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## Calculating the market share via simulation:

- ▶ More detail on step 1:
- ▶ Condition on  $v_i, y_i$  – this is a logit and get closed form
- ▶ Take draws on  $v_i, y_i$  and average over the implied logit shares:

$$\sum_{i=1}^{ns} \frac{\exp(\bar{\mu}_{ij} + \delta_j)}{\sum_k \exp(\bar{\mu}_{ik} + \delta_k)}$$

- ▶ BLP then provide an algorithm (a contraction mapping) that solves for  $\delta$  given the parameters and a set of simulation draws

Having discussed some methods of deriving demand, we now turn to the supply side and a consideration of equilibrium by considering in turn:

- ▶ Estimation of supply side parameters
- ▶ Incorporating multi-product firms
- ▶ Simultaneously estimating supply and demand
- ▶ How to estimate degree of market power or presence of collusion

## Supply Side

- ▶ Simplest models of product differentiation involve single product firms each producing a differentiated product
- ▶ We could begin by specifying a demand system for this set of related products, together with cost functions and an equilibrium notion.
- ▶ The usual assumption is Nash-in-prices.
- ▶ Profits of firm  $j$  are given by

$$\pi_j(p) = p_j q_j(p) - C_j(q_j(p))$$

- ▶ The first order condition is

$$q_j + (p_j - mc_j) \frac{\partial q_j}{\partial p_j} = 0$$

- ▶ We can rewrite as  $p_j = mc_j + b_j(p)$



- ▶ where the price-cost markup is

$$b_j(p) = \frac{q_j}{\left| \frac{\partial q}{\partial p_j} \right|}$$

- ▶ Assume that marginal cost is

$$mc_j = w_j \eta + \lambda q_j + \omega_j$$

- ▶ where  $w_j$  might consist of  $X$  and input prices and  $q$  is output
  - ▶  $\omega_j$  is a supply shock unobserved to the econometrician
- ▶ Combining, the FOC is then

$$p_j = w_j \eta + \lambda q_j + b_j(p) + \omega_j$$

- ▶ If demand parameters are known then the markup is known and can estimate by IV methods (eg 2SLS) where IV are demand-side variables
- ▶ Alternatively mc and demand can be estimated together.

## Multi-Product Firms

- ▶ Non-cooperative oligopolistic Bertrand competition
- ▶ Firm  $f$  produces a subset  $j \in \mathcal{J}_f$  of the products: Profits

$$\sum_{j \in \mathcal{J}_f} (p_j - mc_j) \mathcal{M} s_j(p, X, \xi; \theta)$$

- ▶ where  $\mathcal{M}$  is market size
- ▶  $s_j$  is the simulated aggregate market share
- ▶ Marginal costs

$$mc_j = w_j' \eta + \omega_j$$

- ▶ Any product must have prices that satisfy

$$s_j(p, a) + \sum_{r \in \mathcal{J}_f} (p_r - mc_r) \frac{\partial s_r(p, a)}{\partial p_j} = 0$$

- ▶ Given demand can solve for marginal costs and for  $\omega_j$

- ▶ In vector form, the  $J$  FOC are

$$s - \Omega(p - mc) = 0$$

- ▶ Notice this implies a markup equation  $p - mc = \Omega^{-1}s$
- ▶  $\Omega$  is called the ownership matrix (of dimension  $J \times J$ )
  - ▶ Each element takes on the value of  $\partial s_r(p, a) / \partial p_j$  for every product that the firm owns
- ▶ To estimate the FOC think of estimating the equation

$$mc_j = p_j - b_j(p, x, \xi; \theta) = w_j' \eta + \omega_j$$

- ▶ Just as in estimating demand, estimates of the parameters  $\eta$  can be obtained from orthogonality conditions between  $\omega$  and appropriate instruments

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## Estimation of Supply and Demand Side

- ▶ **Demand side moment:** Restrict the model predictions for product  $j$ 's market share to match the observed market shares

$$S_t^{obs} - s_t(\delta, \theta) = 0$$

- ▶ then solve for the demand side unobservable

$$\xi_{jt} = \delta_{jt}(S, \theta) - x_j' \beta$$

- ▶ **Cost side moment:**
- ▶ Rearranging price FOC's yields

$$mc = p - \Omega^{-1}s$$

- ▶ combined with marginal costs yields cost side unobservable

$$\omega = \ln(p - \Omega^{-1}s) - w' \eta$$

## Nevo: Measuring Market Power in the RTE Cereal Industry

- ▶ The ready-to-eat (RTE) cereal industry is characterized by high price-to-cost margins (PCM) and high concentrations
- ▶ Antitrust authorities accused firms of collusive pricing behavior
- ▶ Nevo tests whether this is the case by estimating the price-cost margin (PCM) and decomposing it into 3 sources:
  - 1 that due to product differentiation
  - 2 that due to multiproduct form pricing and
  - 3 that due to price collusion
- ▶ Overview of methodology:
  - ▶ use the BLP framework to estimate brand-level demand.
  - ▶ use demand estimates and different pricing rules to back out PCMs.
  - ▶ compare PCMs against crude measures of actual PCM to separate the different sources of the markup

## Model and Data

- ▶ Indirect utility is

$$u_{ijt} = \alpha_i p_{jt} + X_j \beta_i + \xi_j + \Delta \xi_{jt} + \varepsilon_{ijt}$$

- ▶ uses brand dummy variables ( $\xi_j$ ) to capture the mean characteristics of RTE cereal
- ▶ once brand dummy variables are included in the regression, the error term is the unobserved city-quarter specific deviation from the overall mean valuation of the brand :structural error is the change in  $\xi_j$  over time (denoted  $\Delta \xi_{jt}$ )
- ▶ Cannot use BLP Type Instruments
  - ▶ there is no variation in each brand's observed characteristics over time and across cities
  - ▶ only variation in IVs from characteristics is due to changes in choice set of available brands
  - ▶ proposes alternative IV to separate the exogenous variation in prices (due to differences in mc) and endogenous variation (due to differences in unobserved valuation)

## IVs with brand dummies

- ▶ Exploit the panel structure of the data (similar to those used by Hausman (1996))
- ▶ The identifying assumption is that, controlling for brand specific means and demographics, city-specific valuations are independent across cities (but are allowed to be correlated within a city)
- ▶ Given this assumption, the prices of the brands in other cities are valid IV's.
  - ▶ prices of brand  $j$  in two cities correlated due to the common  $m_c$
  - ▶ but due to the independence assumption will be uncorrelated with market specific valuation.
  - ▶ One could potentially use prices in all other cities and all quarters as instruments
- ▶ Independence assumption may not hold (for instance, if there is a national demand shock related to health of cereal)



## Identifying Collusive Behavior

- ▶ Recall the markup is given by

$$p - mc = \Omega^{-1}s$$

- ▶ With single product firms the price of each brand is set by a profit-maximizing firm that considers only the profits from that brand. **In this case the ownership matrix will be diagonal**
- ▶ With multi-product firms, firms set the prices of all their products jointly. **In this case some off diagonals will be non-zero**
- ▶ With collusion, firms act as one firm which owns all products (ie joint profit-maximization of all the brands). **In this case the ownership matrix will have no zeros**
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## Results

- ▶ Compares predicted PCM under all three situations to the observed PCM calculated using accounting data for costs
- ▶ Finds that the first two effects explain most of the observed price-cost margins
- ▶ Prices in the industry are consistent with noncollusive pricing behavior, despite the high price-cost margins.